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# **Evaluation of re-analyses over China based on the temporal asymmetry of daily temperature variability**

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#### **Abstract**

As an intrinsic feature of daily surface air temperature (SAT) variability found in station measurements, temporal asymmetry (TA) can be taken as an evaluation metric to access the quality of SAT re-analysis product. In this study, TA calculated from four SAT variables, i.e., daily mean SAT (T<sub>mean)</sub>, daily maximum SAT (T<sub>max</sub>), daily minimum SAT (T<sub>min</sub>) and diurnal temperature range ( $T_{DTR} = T_{max} - T_{min}$ ), is applied to evaluate synoptic-scale performance of four reanalysis products (NCEP-2, JRA-55, ERA-I, and ERA-5) over China. The results show that four re-analyses overall overestimate the TA of daily  $T_{max}$  and  $T_{min}$  variability over China, but with a comparatively consistent estimated TA for  $T_{mean}$ . Moreover, the TA of  $T_{mean}$  variability for these four re-analyses shares high spatial consistency with those from the observation. However, four re-analyses own the similar region-dependent spatial patterns of overestimated TA for  $T_{max}$  and  $T_{min}$  variability, especially for  $T_{max}$ . Since high TA is an indicator for strong nonlinear feature, only  $T_{mean}$  reanalysis is the most suitable to explore synoptic-scale extreme events, such as heat waves and cold waves, which are highly related to the strong nonlinear processes.

### **1 Introduction**

Asymmetric phenomena are ubiquitous in both natural and social sciences (Heinrich [2004;](#page-11-0) King [1996;](#page-11-1) Livina et al. [2003;](#page-11-2) Ashkenazy and Tziperman [2004;](#page-10-0) Lisiecki and Raymo [2005;](#page-11-3) Bartos and Jánosi [2005](#page-11-4); Gyure et al. [2007](#page-11-5); Ashkenazy et al. [2008,](#page-11-6) [2016;](#page-11-7) Bisgaard and Kulahci [2011;](#page-11-8) Xie et al. [2016,](#page-12-0) [2019\)](#page-12-1), and it is an important indicator of nonlinear underlying processes (Schreiber and Schmitz [2000;](#page-11-9) Bartos and Jánosi [2005](#page-11-4); Gyüre et al. 2007; Ashkenazy et al. [2008](#page-11-6); Roldan and Parrondo [2010](#page-11-10); Lacasa et al. [2012](#page-11-11)).

#### **Key points**

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As a kind of asymmetry, the temporal asymmetry (TA) in time series, defned by diferent statistics between forward and backward (reversed) directed series, plays an important role in air temperature variability studies (Bartos and Jánosi [2005](#page-11-4); Gyüre et al. 2007; Ashkenazy et al. [2008](#page-11-6); Xie et al. [2016,](#page-12-0) [2019](#page-12-1)). Previous studies also found that there exists diferential TA among diferent temperature variables' daily fuctuations over China from both station observations and NCEP-2 re-analyses (Xie et al. [2019](#page-12-1)). For commonly used temperature variables, daily mean 2-m surface air temperature (SAT,  $T_{mean}$ ), daily maximum SAT ( $T_{max}$ ), daily minimum SAT  $(T_{min})$ , and diurnal temperature range  $(T_{\text{DTR}} = T_{\text{max}} - T_{\text{min}})$ , TA strengths among these four temperature variables are markedly diferent with the weakest TA for DTR and the strongest TA for  $T_{mean}$ . Compared with the TA from station observations, TA from  $T_{\text{max}}$  and  $T_{\text{min}}$ in National Centers for Environmental Prediction (NCEP) AMIP-II Reanalysis (hereafter NCEP-2, Kanamitsu et al. [2002](#page-11-12)) is highly overestimated for most regions over China (Xie et al. [2019\)](#page-12-1).

TA has been found to be an intrinsic feature in daily SAT variability in station measurements (Bartos and Jánosi [2005](#page-11-4); Gyüre et al. 2007; Xie et al. [2016,](#page-12-0) [2019](#page-12-1); Li et al. [2021](#page-11-13)). Although similar TA behavior was also reported in re-analysis data (Ashkenazy et al. [2008;](#page-11-6) Xie et al. [2019](#page-12-1); Li et al. [2021\)](#page-11-13), only a few studies (Xie et al. [2019;](#page-12-1) Li et al. [2021\)](#page-11-13)

<sup>•</sup> Four re-analyses all overestimate the temporal asymmetry of daily maximum and minimum temperature variability over China.

<sup>•</sup> For overestimated temporal asymmetry, four re-analyses own the similar region-dependent spatial patterns.

<sup>•</sup> Daily mean air temperature variability of re-analyses is the most suitable for extreme event study.

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compared TA diference between station measurement and re-analysis products, detailed and systematic comparative studies are still required. More specifcally, is it a universal feature to diferent kinds of re-analyses for the reported overestimated TA in  $T_{max}$  and  $T_{min}$  variability from NCEP-2 re-analyses over China (Xie et al. [2019](#page-12-1); Li et al. [2021\)](#page-11-13)? If the answer is yes, is there any spatial consistency among diferent reanalysis products?

The reanalyzed products are outputs from the assimilation technology of numerical weather prediction to restore the observation data and compensate for the lack of uneven distribution of weather stations (Bengtsson and Shukla [1988](#page-11-14); Trenberth and Olson [1988;](#page-11-15) Chen and Liu [2016\)](#page-11-16). Previous studies used diferent methods to compare and evaluate reanalysis data at diferent temporal and spatial scales from the point of view of mean value, standard deviation, long-term trend, long-range correlation (LRC), and so on, and they found re-analyses do not always work well to reproduce consistent results with observations (Flocas et al. [2005;](#page-11-17) Ma et al. [2008](#page-11-18); Pitman and Perkins [2009;](#page-11-19) Mao et al. [2010](#page-11-20); Marques et al. [2010;](#page-11-21) Mooney et al. [2011;](#page-11-22) Alfred et al. [2011;](#page-10-1) You et al. [2011,](#page-12-2) [2013;](#page-12-3) Cornes and Jones [2013;](#page-11-23) Chen and Iwasaki [2014;](#page-11-24) Taguchi [2017](#page-11-25); Zhu et al. [2017;](#page-12-4) He et al. 2018; Alghamdi [2020](#page-10-2)). For example, He and Zhao ([2018](#page-11-26)) evaluated the air temperature reanalysis variables by means of LRC, and they found that NCEP-2 reanalysis overestimates LRC over some specifc regions. Zhu et al. [\(2017](#page-12-4)) revealed that Interim European Centre for Medium-Range Weather Forecasts (ECMWF) Re-Analysis (ERA-Interim, herein ERA-I, Dee and Coauthors [2011\)](#page-11-27) can capture the intensity indices of the continuous extreme temperature events, but not their frequency indices. Since TA is closely related to the synoptic-scale processes (Ashkenazy et al. [2008](#page-11-6); Xie et al. [2016;](#page-12-0) Li et al. [2021](#page-11-13); Quan et al. [2021](#page-11-28)) and nonlinear behaviors (Li et al. [2021](#page-11-13)), behaviors like extreme temperature events found in ERA-I should also be able to be revealed by TA.

Besides NCEP-2 reanalysis and ERA-I, other commonly used reanalysis products are available: the 55-year Japanese Project (hereafter JRA-55, Ayataka et al. [2011;](#page-11-29) Kobayashi et al. [2015](#page-11-30)), and the Fifth generation ECMWF Re-Analysis (hereafter ERA-5, Radu et al. [2018;](#page-11-31) Hersbach et al. [2020](#page-11-32)). These four reanalysis products will be exploited to test whether there are universal conclusions on TA related to SAT variability.

Moreover, Ye and Hsieh [\(2008](#page-12-5)) found that increasing nonlinearity in ENSO and Lorenz systems can enhance their predictability by improving the contributions from the low-frequency variations. The process-dependent intrinsic predictability is also reported to be enhanced with increased nonlinearity (Huang and Fu [2019](#page-11-33)). As one kind of nonlinearity, overestimated TA in the NCEP-2 reanalysis daily minimum and maximum air temperature anomaly series over China is concurrent with overestimated intrinsic predictability, prediction skill, and the occurrence number of extreme events (Li et al. [2021](#page-11-13)). This fnding indicates that overestimated TA in reanalysis products may distort the conclusions on extreme event studies. By comparing the TA calculated from the station observations of diferent SAT variables to those from diferent reanalysis products, the most suitable SAT variable and the corresponding reanalysis products to estimate TA can be determined.

The paper is organized as follows. Section [2](#page-1-0) summarizes the sources of observation and reanalysis data and the methods used in this study. Section [3](#page-3-0) reports the overestimated TA in both  $T_{\text{max}}$  and  $T_{\text{min}}$  from observations and re-analyses, contrasting to the comparable consistency in TA for  $T_{\text{mean}}$ . And quantitative comparison of TA between observations and re-analyses for four diferent SAT variables is also provided. At last, Section [4](#page-9-0) concludes this study with a brief discussion.

## <span id="page-1-0"></span>**2 Data and methods**

#### **2.1 Data**

#### **2.1.1 Observations**

The observational time series (hereafter observations) of  $T_{\rm{mean}},$   $T_{\rm{max}},$   $T_{\rm{min}},$  and  $T_{\rm{DTR}}$  were downloaded from China Meteorological Administration. After data quality check by removing the time series with missing points, datasets of 643 Chinese meteorological stations were fnally selected for this study, covering a total of 40 years from 1979 to 2018.

#### **2.1.2 Reanalysis**

For the reanalysis data of 2-m air temperature from 1979 to 2018, we selected four commonly used products: NCEP-2 (Kanamitsu et al. [2002](#page-11-12)), JRA-55 (Ayataka et al. [2011](#page-11-29); Kobayashi et al. [2015\)](#page-11-30), ERA-Interim (Dee and Coauthors [2011](#page-11-27)) and ERA-5 (Radu et al. [2018](#page-11-31); Hersbach et al. [2020](#page-11-32)). NCEP-2 products were acquired from the National Oceanic and Atmospheric Administration (NOAA) and its website at <http://www.esrl.noaa.gov/psd>. JRA-55 is produced by the Japan Meteorological Agency (JMA) [\(https://climatedatagu](https://climatedataguide.ucar.edu/climate-data/jra-55)[ide.ucar.edu/climate-data/jra-55\)](https://climatedataguide.ucar.edu/climate-data/jra-55). From the ECMWF's website: [https://www.ecmwf.int/en/forecasts/datasets/browse](https://www.ecmwf.int/en/forecasts/datasets/browse-reanalysis-datasets)[reanalysis-datasets,](https://www.ecmwf.int/en/forecasts/datasets/browse-reanalysis-datasets) both ERA-I and ERA-5 are downloaded. Detailed information about these four reanalysis datasets is shown in Table [1](#page-2-0).

Previous studies found that topography, altitude, and sparse station are the main factors afecting the quality of reanalysis data (Rusticucci and Kousky [2002;](#page-11-34) Wang et al. [2015;](#page-11-35) Zhao et al. [2018;](#page-12-6) He and Zhao [2018\)](#page-11-26). Due to the

<span id="page-2-0"></span>**Table 1** Details of the four reanalysis datasets used in this study

Dataset	Horizontal resolution	Output frequency	Assimilation method	Source
NCEP-2	$1.875^{\circ} \times 1.904^{\circ}$	6-hourly for analyses	3D-Var	<b>NOAA</b>
$JRA-55$	$1.25^{\circ} \times 1.25^{\circ}$	6-hourly for analyses	4D-Var	JMA
ERA-I	$0.75^{\circ} \times 0.75^{\circ}$	6-hourly for analyses	4D-Var	<b>ECMWF</b>
ERA-5	$0.5^\circ \times 0.5^\circ$	Hourly throughout	4D-Var ensemble	<b>ECMWF</b>

*3D-Var* or *4D-Var* 3-dimensional or 4-dimensional variational data assimilation.

sparse stations in Tibetan region, we removed the stations over Tibetan regions and only evaluated the quality of the reanalysis data by using observations from the rest of the stations over China except Tibetan regions.

In this study, all referred time series are standardized by removing the annual cycle (Koscielny-Bunde et al. [1998\)](#page-11-36) by  $T_i' = T_i - \langle T_i \rangle$ ,  $i = 1, ..., N$ , where  $T_i$  is any given daily air temperature variables,  $\langle T_i \rangle$  is its long-time average for each calendar day,  $T_i'$  is the corresponding anomaly, and  $N$  is the data length.

#### **2.2 Methods**

#### **2.2.1 Interpolation for reanalysis**

In order to compare TA measures calculated from the four reanalysis products with diferent spatial resolutions to those from station observations, we interpolated the two-dimensional gridded data into the corresponding observation stations (hereafter interpolated reanalysis). Two widely used methods in previous studies were adopted: Gaussian weight function (Maddox et al. 1981; Xie et al. [2019](#page-12-1)) and interpolation by directly choosing the closest grid point to represent the targeted station (Diaconescu et al. [2018](#page-11-37); Pendergrass and Knutti [2018;](#page-11-38) Yang et al. [2020](#page-12-7)).

For Gaussian weight function, the interpolated reanalysis is given by

$$
\text{Re}_s = \frac{\sum_i \sum_j w(i,j)s(i,j)}{\sum_i \sum_j w(i,j)},\tag{1}
$$

where  $s(i, j)$  is the targeted variable over grid  $(i, j)$  and the weight function  $w(i, j)$  is:

$$
w(i,j) = \exp\left(\frac{-d^2(i,j,k)}{4C^2}\right),\tag{2}
$$

where *C* is the weighting constant, the Euler distance *d* is from the grid point  $(i, j)$  to the location of the specific station. For the reanalyzed data with diferent resolutions, detailed numerical calculations found that the smaller value *C* (for example  $C = 0.5$  or  $C = 1.0$ ) within a certain range can produce reliable interpolations.

#### **2.2.2 Measuring temporal asymmetry of time series**

Diferent methods have been proposed and applied to measure the TA strength of air temperature variations (Xie et al. [2016](#page-12-0), [2019;](#page-12-1) Zhang et al. [2019](#page-12-8); Li et al. [2021\)](#page-11-13). Zhang et al. [\(2019](#page-12-8)) compared several TA measure methods and found all of them perform nearly equally well for most of cases. For simplicity, the asymmetry index (*A*) is adopted to quantify the TA strength in the air temperature anomaly  $T'_{i}$ , which is defned as the ratio of positive air temperature variability steps to the total (positive plus negative) steps (Ashkenazy et al. [2008](#page-11-6)):

<span id="page-2-1"></span>
$$
A = \frac{p}{p+n} = \frac{\sum_{i=1}^{i=N-1} \theta(T'_{i+1} - T'_i)}{\sum_{i=1}^{i=N-1} \theta(T'_{i+1} - T'_i) + \sum_{i=1}^{i=N-1} \theta(-(T'_{i+1} - T'_i))},\tag{3}
$$

where  $\theta(x) = \begin{cases} 1, & x > 0 \\ 0, & x < 0 \end{cases}$ , *N* is the length of *T'*<sub>*i*</sub>. When the value of TA is smaller than  $0.5$   $(A < 0.5)$ , it indicates that the air temperature warms rapidly and gradually becomes cold. The value of TA is larger than  $0.5 (A > 0.5)$  indicates that the air temperature rapidly cools and gradually warms. When the value of TA is close to 0.5 ( $A \approx 0.5$ ), the air temperature time series are symmetric.

#### **2.2.3 Signifcance test**

To carry out the signifcance tests, the iterative amplitude Fourier transform (IAAFT) is employed to generate surrogates for each time series by keeping the two-point correlation and probability density function as those from the original series. Taken  $T_{mean}$  anomaly time series as an example, we generate 500 surrogate series for it, and calculate *A* by means of Eq. [\(3](#page-2-1)) for each surrogate to obtain 500 values for *A*. Sorting these 500 values for *A* can defne the lower 1% levels and the upper 99% among these 500 values of *A* as the signifcance thresholds. And then two thresholds of *Ac*<sup>1</sup> and *Ac*2 can be obtained. When the calculated value of *A* for a given series is lower than  $A_{c1}$  or higher than  $A_{c2}$ , then this given series is taken to be statistically signifcantly temporal asymmetric. Such signifcance test through IAAFT surrogate can ensure that our results are not artifcially infuenced by

stochastic effects, the autocorrelation and probability distribution of a time series (Huang et al. [2020](#page-11-39)).

## <span id="page-3-0"></span>**3 Results**

## **3.1 Overestimated temporal asymmetry in T**<sub>max</sub> **and Tmin**

First of all, the estimated temporal asymmetry (TA) for four SAT variables' variability from the observations is compared with those from four kinds of re-analyses, and detailed results are presented in Fig. [1.](#page-3-1) Here, the determination of overestimation or underestimation should be based on the statistical signifcance test. For each estimated TA, there is an interval from the signifcance test to judge whether it is overestimated or underestimated. The critical values at the signifcance level of 0.02 are shown as the lines on both sides of the 1:1 line in Fig. [1.](#page-3-1) Only outside of the critical lines, the estimated TA can be taken as signifcantly overestimated or underestimated at the signifcance level of 0.02. The marked results are that nearly all values of estimated TA in  $T_{\text{max}}$  and  $T_{\text{min}}$  from all four re-analyses are overestimated (see Fig. [1b, c](#page-3-1)), especially in  $T_{\text{max}}$ , only a few from NCEP-2 reanalysis are comparable with or lower than those from observations. This fnding is consistent with previous studies in NCEP-2 reanalysis compared with limited station observations (Xie et al. [2019\)](#page-12-1). Diferent from the results given by only one specifc reanalysis product (NCEP-2), the results given here indicate that the overestimated TA in  $T_{\text{max}}$  and  $T_{\text{min}}$  from re-analyses may be taken as a common intrinsic feature to all analyzed reanalysis products.

Secondly, overestimated TA in both  $T_{\text{max}}$  and  $T_{\text{min}}$  from all four kinds of re-analyses has its well-defned spatial



<span id="page-3-1"></span>**Fig. 1** Temporal asymmetry scatter plots between observations and interpolated re-analyses for **a**  $T_{\text{mean}}$ , **b**  $T_{\text{max}}$ , **c**  $T_{\text{min}}$ , **d**  $T_{\text{DTR}}$  from 1979 to 2018. The two horizontal and vertical dash black lines denote the two critical thresholds for the confdence interval of *A* at the signif-

cance level of 0.02, respectively. The dash cyan line denotes the 1:1 diagonal line. The dash lines on both sides of the 1:1 line are the thresholds at a signifcance level of 0.02

patterns (see Figs. [2](#page-4-0) and [3](#page-5-0)). Especially for  $T_{\text{max}}$ , four kinds of re-analyses share very similar spatial distribution of TA diference between observation and interpolated reanalysis, which is defined as  $A_{\text{diff}} = A_{\text{re}} - A_{\text{obs}}$  with  $A_{\text{re}}$  from reanalysis and *A*<sub>obs</sub> from observations. There are region-dependent spatial patterns of  $A_{\text{diff}}$ . Weak overestimated TA occupies most of the regions of China. However, there are still some special patterns. The consistent estimation of TA between observation and interpolated reanalysis mainly locates over northeast part of Heilongjiang Province and west part of both Yunnan and Sichuan Provinces. Among four kinds of re-analyses, the shared patterns of the strongest overestimated TA lie in Guizhou and Chongqing provinces. These regions are also where the TA is not statistically signifcant in  $T_{\text{max}}$  for observations (see Fig. [4f](#page-6-0)), but strong TA for four kinds of re-analyses (see Fig.  $4g-j$ ). According to studies on temperature asymmetry and the mechanism of temperature asymmetry (Ashkenazy et al. [2008](#page-11-6); Piskala and Huth [2020](#page-11-40); Quan et al. [2021](#page-11-28)), the possible cause behind regiondependent spatial patterns of  $A_{\text{diff}}$  is closely related to frontal processes, and frontal processes have a strong regional dependence (Ashkenazy et al. [2008](#page-11-6); Piskala and Huth [2020](#page-11-40); Quan et al. [2021\)](#page-11-28). This may be part of the reason for TA's regional dependence, and the mechanism behind this phenomenon deserves further study in depth.

For  $T_{\text{min}}$ , no common national-scale spatial pattern of TA diference is shared among four kinds of re-analyses between observation and interpolated reanalysis over China (see Fig. [3](#page-5-0)). Only NCEP-2 and ERA-5 share the similar national-scale spatial distribution of overestimated TA (see Fig. [3a, d\)](#page-5-0). However, only JRA-55 and ERA-I rather than NCEP-2 and ERA-5 share the similar spatial distribution of TA diference with an east-west dipole pattern. Over the half part of China east to 110° E, nearly all values of TA are overestimated, whereas over the another half part of China west to 100° E, nearly all values of TA are weakly underestimated. The contrasting spatial patterns of TA in  $T_{\text{min}}$  are mainly due to the weak TA strength in observations (see Fig. [5a](#page-7-0)) but strong TA strength in Central parts of China in four kinds of re-analyses (see Fig. [5b–e\)](#page-7-0). At the same time, contrast topography and possible related specifc procedures adopted by JRA-55 and ERA-I reanalysis products (Pinheiro et al. [2020\)](#page-11-41) may contribute to TA diference with an eastwest dipole pattern in  $T_{\text{min}}$ .

Besides the almost identical patterns of TA diference for  $T_{\text{min}}$ , JRA-55, and ERA-I also share almost the same TA patterns for all temperature variables ( $T_{mean}$ ,  $T_{max}$ ,  $T_{min}$ , and  $T<sub>DTR</sub>$ ), details can be found in Fig. [4c, d, h, and j](#page-6-0); Fig. [5c,](#page-7-0) [d, h, and j.](#page-7-0) The similarity between ERA-I and JRA-55 reanalyses on synoptic-scale phenomena has been reported in



<span id="page-4-0"></span>**Fig. 2** Spatial distribution of *A* difference in  $T_{\text{max}}$  from 1979 to 2018 between observation and interpolated reanalysis for **a** NCEP-2, **b** JRA-55, **c** ERA-I, and **d** ERA-5, the black asterisk indicates the con-

fdence interval for the same TA estimation between observation and interpolated reanalysis at the signifcance level of 0.02



<span id="page-5-0"></span>**Fig. 3** Same as Fig. [2](#page-4-0) but for  $T_{min}$ 

the literature. When Pinheiro et al. [\(2020\)](#page-11-41) studied subtropical cut-off lows in the southern hemisphere, they found that the diferences of track density between ERA-I and JRA-55 are relatively small. They explained this similarity between ERA-I and JRA-55, as both of them used the same data assimilation systems (Pinheiro et al. [2020\)](#page-11-41).

## **3.2 Comparable estimation of temporal asymmetry in T**<sub>mean</sub>

Different from the markedly overestimated TA found in  $T_{\text{max}}$ and  $T_{\text{min}}$ , more consistent TA estimations are revealed in  $T_{mean}$  between observations and re-analyses (see Fig. [6](#page-8-0)). There are only a few stations (less than 1%) with strong overestimated TA, and nearly 50% of the stations have the same TA estimation for both observations and re-analyses from NCEP-2 and ERA-5 (see Fig. [6a, d](#page-8-0)). Moreover, nearly all values of estimated TA in  $T_{mean}$  from both observations and re-analyses are statistically signifcant at the signifcance level of 0.02 (only less than 2% stations in NCEP-2 have insignifcant TA estimation). For clear view, the same critical thresholds based on the signifcance test from the surrogated linear process (see details in 2.2.3 Signifcance test) are marked with vertical and horizontal black dotted lines in the scatter plot, as shown in Fig. [1](#page-3-1). The lower value of the critical threshold  $A_{c1}$  is 0.4925 and the upper value  $A_{c2}$ 0.5074 (vertical and horizontal black dotted lines in Fig. [1a](#page-3-1)).

Since nearly all values of *A* for the reanalysis products are all larger than these critical values (see Fig. [1a\)](#page-3-1), nearly all the daily mean SAT variability takes the temporal asymmetric behaviors over China.

The regions of the strongest TA strength lie over central and southeast regions of China, especially to the south of Qinling Mountains and to east of 105° E (see Fig. [4a](#page-6-0)). Over these regions, the variations of  $T_{mean}$  gradually warm and rapidly cool. Such a temporal asymmetry phenomenon occurring at mid-latitudes has been attributed to the diferent contributions from the warm and cold fronts (Ashkenazy et al. [2008](#page-11-6); Piskala and Huth [2020](#page-11-40); Quan et al. [2021\)](#page-11-28). Lower *A* is found in the Basin of Tarim, Qaidam and Sichuan, Yun-Gui Plateau and northeast of China (see Fig. [4a](#page-6-0)). Surprisingly, all four re-analyses can reproduce the high and low values of TA over these regions (see Fig. [6\)](#page-8-0), and this is why there are more consistent TA estimations in  $T_{mean}$  (see Fig. [1a](#page-3-1)).

#### **3.3 TA spatial pattern similarity quantifcation**

In order to quantify the TA spatial pattern similarity between observations and interpolated re-analyses, we adopt the Tay-lor diagram (Taylor [2001\)](#page-11-42) to compare it quantitatively. Three statistics most often used to quantify pattern similarity in Taylor diagram are the correlation coefficient between observations and interpolated re-analyses, defned as



<span id="page-6-0"></span>**Fig. 4** Spatial distribution of *A* in  $T_{mean}$  (left column) and  $T_{max}$  (right column) from 1979 to 2018 between observation (**a**, **f**) and interpolated reanalysis: NCEP-2 (**b**, **g**), JRA-55 (**c**, **h**), ERA-I (**d**, **i**), and

ERA-5 (**e**, **j**). The black asterisk indicates the confdence interval with two critical thresholds  $A_{C1} = 0.4925$  and  $A_{c2} = 0.5074$  for TA estimation at the signifcance level of 0.02

$$
RCC = \frac{\sum_{i=1}^{n} \left( A_{Ra} - \overline{A_{Ra}} \right) \left( A_o - \overline{A_o} \right)}{n \sigma_{A_{Ra}} \sigma_{A_o}},
$$
\n(4)

the root-mean-square (RMS) diference between observations and interpolated re-analyses standardized by the results from observations, defned as

$$
SRMS = \sqrt{\frac{\sum_{i=1}^{n} \left[ \left( A_{Ra} - \overline{A_{Ra}} \right) - \left( A_o - \overline{A_o} \right) \right]^2}{n\sigma_{A_o}^2}},
$$
(5)

and the standard deviation of  $A_{Ra}$  standardized by  $A_{o}$ 



<span id="page-7-0"></span>**Fig. 5** Same as Fig. [4](#page-6-0) but for  $T_{min}$  (left column) and  $T_{DTR}$  (right column)

$$
SSTD = \frac{\sigma_{A_{Ra}}}{\sigma_{A_o}},\tag{6}
$$

with *n* the number of stations,  $\sigma_{A_{\rho}}$  and  $\sigma_{A_{R_a}}$  are standard deviation for *A* from observations and a specifc interpolated reanalysis. For perfectly consistent TA estimations between observations and interpolated reanalysis over all considered regions over China, *RCC*=1, *SRMS*=0 , and *SSTD*=1.

Taylor diagrams constructed from four kinds of interpolated re-analysis and observations for TA of four SAT variability could more accurately evaluate the quality of diferent reanalysis products for diferent SAT variables (Fig. [7](#page-8-1)). First of all, it should be pointed out that synoptic-scale performance of all analyzed re-analyses is not comparable to their climate-scale performance. For  $T_{mean}$ , the four reanalysis products all perform the best among these four



<span id="page-8-0"></span>**Fig. 6** Same as Fig. [2](#page-4-0) but for  $T_{mean}$ 

<span id="page-8-1"></span>**Fig. 7** Taylor diagram for **a**  $T_{\text{mean}}$ , **b**  $T_{\text{max}}$ , **c**  $T_{\text{min}}$ , **d**  $T_{\text{DTR}}$ of *A* from observation and interpolated reanalysis. The black dot stands for the results calculated from observations, which serves as the reference. The radial distance from the origin is proportional to the standard deviation of a pattern normalized by reference pattern, the centered root mean square (RMS) diference between the reference and re-analyses is proportional to their distance apart and the correlation between the reference and re-analyses is given by the azimuthal position of a given reanalysis



SAT variables, and they are all located in a limited region in the Taylor diagram (see Fig. [7a\)](#page-8-1). Especially ERA-I, ERA-5, and JRA-55 are nearly of the equally well performance with the highest RCC value (around 0.8) and the lowest SRMS value (around 0.62), only NCEP-2 a little bit worse with the RCC value around 0.73 and SRMS value around 0.78. Four reanalysis products perform second well for  $T_{\text{max}}$  among these four SAT variables (see Fig. [7b](#page-8-1)), among them, ERA-5 is the best with the RCC value around 0.73 and SRMS value around 0.74 and NCEP-2 the worst with the RCC value around 0.4 and SRMS value around 1.14. There are great discrepancies in TA estimation between the observation and the interpolated re-analyses for both  $T_{\text{min}}$  and  $T_{\text{DTR}}$  (Fig. [7c, d](#page-8-1)), for all re-analyses, the RCC values are low (less than 0.5), and SRMS values are all high (larger than 0.99).

Moreover, Ye and Hsieh found (2008) that increasing nonlinearity in ENSO and Lorenz systems can enhance their predictability by improving the contributions from the low-frequency variations. As a kind of nonlinearity, temporal asymmetry in daily SAT variability is closely related to extreme events and some small-scale phenomena (Raghavendra et al. [2018](#page-11-43); Li et al. [2021\)](#page-11-13). More comparable TA estimation between observations and re-analyses leads daily mean surface temperature reanalysis products to be the most suitable choice to synoptic-scale extreme event study.

## <span id="page-9-0"></span>**4 Discussion and conclusion**

In order to directly compare the TA from observation with those from reanalysis, the grid reanalysis data have been interpolated into the targeted station to reach the interpolated reanalysis data. In previous studies, a number of interpolation methods were proposed and applied to test the impacts of interpolations on the derived results (Maddox et al. 1981; Xie et al. [2019](#page-12-1)). Xie et al. [\(2019](#page-12-1)) found that if the suitable choice is made to the interpolation distance parameter *C*, interpolations do not change the TA calculations too much. We compared the effects from different methods and different choices of interpolation distance parameter *C* on the calculations of TA from the original and four interpolated reanalysis data, detailed results can be found in Fig. [8](#page-9-1) and were summarized in Table [2.](#page-10-3) It is confrmed that if the suitable distance parameter is chosen, the estimation of TA is insensitive to the interpolations. Also the diferent interpolation methods do not change the TA calculations too much. Especially, the estimation of TA from ERA-I and JRA-55





<span id="page-9-1"></span>**Fig.** 8 Probability density function (PDF) for *A* of  $T_{\text{mean}}$  from **a** NCEP-2, **b** JRA-55, **c** ERA-I, and **d** ERA-5 by diferent interpolation ways. Gau means the Gaussian weight function interpolation with

different values of  $C$  ( $C=1.5$ , green;  $C=1.0$ , cyan;  $C=0.5$ , blue), Method-2 represents interpolation by the closest points to the stations (red) and Grid-ori from original reanalysis (gray) shadow

<span id="page-10-3"></span>

is more robust to the interpolation methods and choices of distance parameter (Fig. [8](#page-9-1); Table [2\)](#page-10-3).

As an important nonlinear indicator, temporal asymmetry of time series can be taken as an intrinsic feature of nonlinear time series. Similar results can be found in a simple nonlinear model, such Logistic map with a given regime (Figure not shown here). TA calculation from observation in daily mean surface air temperature variability over China, (Fig. [1a\)](#page-3-1), having nearly totally statistically signifcant temporal asymmetry leads it to be an intrinsic feature of daily surface air temperature variability. It can be taken as a metric to access the quality of diferent daily SAT reanalysis products. Taking TA as an evaluation measure, four daily SAT reanalysis products (NCEP-2, ERA-I, ERA-5, JRA-55) are accessed. Compared with the observations, the four reanalyses can consistently reproduce the TA in  $T_{\text{mean}}$  over South China. NCEP-2 underestimates the TA in  $T_{mean}$  over the northeast and central regions of China. JRA-55, ERA-I, and ERA-5 overestimate the TA in  $T_{mean}$  over the northwest and central regions. However, all four re-analyses universally overestimate the TA in  $T_{\text{max}}$  and  $T_{\text{min}}$ . These results are different from the fndings by previous studies based on the linear view (Zhao et al. [2018](#page-12-6); He and Zhao [2018\)](#page-11-26). Due to the weak TA in the  $T_{\text{DTR}}$  (see Figs. [1d](#page-3-1) and [5f–j\)](#page-7-0), there are larger relative uncertainties in calculation and comparison of TA from the observations and reanalysis, so no conclusive results can be reached on  $T_{\text{DTR}}$ .

It was reported that the modeled daily maximum temperature and daily minimum temperature are unsuitable for the study of extreme events such as heat waves due to the underestimated daily maximum or overestimated daily minimum temperature compared with observations (Raghavendra et al. [2018](#page-11-43)). Based on the results of TA estimation from all four reanalysis products, we confrm that the nonlinear strength is highly overestimated in  $T_{\text{max}}$  and  $T_{\text{min}}$  from the re-analyses. Since there is a close relation between nonlinearity and extreme events (Ye and Hsieh [2008;](#page-12-5) Li et al. [2021](#page-11-13)), results from daily maximum or minimum temperature reanalyses may distort conclusions on extreme events such as heat waves and cold waves. On the contrary, the most comparable consistency of TA estimation in daily mean surface air temperature variability from both observations and re-analyses make it to be a reasonable choice. Last but not least, detailed comparison among the TA results for seasonal (Figure not shown here) and all-year data (see Fig. [1](#page-3-1)) shows that all these conclusions given above are qualitatively similar with only minor quantitative diferences among results for four seasons, i.e., the above conclusions are not season-dependent.

**Code availability** Not applicable.

**Author contribution** Study design: Zuntao Fu, Wenmi Chai; data collection: Wenmi Chai; statistical analysis: Wenmi Chai; result interpretation: Zuntao Fu, Wenmi Chai; manuscript preparation, and review: Wenmi Chai, Yu Huang, Lichao Yang, Heng Quan, and Zuntao Fu.

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**Data availability** All observed records for these four temperature variables were downloaded from the China meteorological data sharing service system [\(http://cdc.cma.gov.cn](http://cdc.cma.gov.cn)). Reanalysis NCEP-2 products were acquired from the National Oceanic and Atmospheric Administration (NOAA) and its website at [http://www.esrl.noaa.gov/psd.](http://www.esrl.noaa.gov/psd) JRA-55 is produced by the Japan Meteorological Agency (JMA) [\(https://clima](https://climatedataguide.ucar.edu/climate-data/jra-55) [tedataguide.ucar.edu/climate-data/jra-55\)](https://climatedataguide.ucar.edu/climate-data/jra-55). From the ECMWF's website: [https://www.ecmwf.int/en/forecasts/datasets/browse-reanalysis](https://www.ecmwf.int/en/forecasts/datasets/browse-reanalysis-datasets)[datasets](https://www.ecmwf.int/en/forecasts/datasets/browse-reanalysis-datasets), both ERA-I, and ERA-5 are downloaded.

#### **Declarations**

**Ethics approval** We confrm that this article is an original research and has not been published or presented previously in any journal or conference in any language (in whole or in part).

**Consent to participate and consent for publication** We have consent to participate and publish.

**Conflict of interest** The authors declare no competing interests.

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